

# MADRAS SCHOOL OF ECONOMICS

UNDERGRADUATE PROGRAMME IN ECONOMICS (HONOURS) [2023-26]

SEMESTER 6 [JANUARY – APRIL, 2026]

REGULAR EXAMINATION, APRIL-MAY 2026

Course Name: Introduction to Time Series Analysis Course Code: DE09

Duration: 2 Hours

Maximum Marks: 60

**Instructions:** For part A write short answers (most preferably a single word,/single number/single sentence). For part B, writing relevant formulae and quoting correct definitions helps me give you part marks. For questions with qualitative answers, creative writing is prohibited.

**Part A: Answer all questions in one line.** (1 mark x 10 questions = 10 marks)

Notations are as used in class.

1. What is the full form of PACF?
2. Define an integrated process of order  $n$ , written as  $I(n)$ .
3. State one implication of asymptotic normality of MLE for inference.
4. Write down the formula for Akaike Information Criteria. Briefly mention the meaning of the terms.
5. For a first order deterministic sequences  $y_t = 0.3y_{t-1} + \epsilon_t$ , write the impulse response function  $\psi_j$  as a function of  $j$ .
6. If  $a_t + b_t = c_t$  and all three are  $I(1)$ , give a cointegrating vector for  $(a, b, c)^T$ .
7. What is the null hypothesis of the Ljung–Box test? You can describe it in words.
8. For  $Y_t = 0.5Y_{t-1} + \epsilon_t$ , what is  $E[Y_{t+1} | \mathcal{F}_t]$ ?
9. Let  $W(t)$  be a Wiener process. Compute  $\mathbb{E} \left( \int_0^1 W(t)^2 dt \right)$ .
10. The process  $Y_t = 1.5Y_{t-1} - 0.5Y_{t-2} + \epsilon_t$  is an ARIMA( $p, d, q$ ) process. What are the values of  $p, d, q$ ?

**Part B: Answer any five.** (5 mark x 10 questions = 50 marks)

11. [10 marks] What is Granger causality? Explain how VAR models are used in detecting Granger causality. Illustrate the use with a real example.
12. [10 marks] State and prove the Wiener-Kolmogorov theorem. You may assume that the optimal MMSE forecast is conditional expectation.
13. Suppose a covariance-stationary process has sample autocorrelations  $\hat{\rho}_1 = 0.6$  and  $\hat{\rho}_2 = 0.2$ .
  - (a) [4 marks] Assuming the process is an AR(2) model, write down the Yule-Walker equations in matrix form.
  - (b) [4 marks] Solve the system to obtain the method-of-moments estimators  $\hat{\phi}_1$  and  $\hat{\phi}_2$ .
  - (c) [2 marks] Verify that the estimated AR(2) process is covariance stationary.
14. A zero-mean covariance-stationary process  $\{Y_t\}$  has autocovariances

$$\gamma_0 = 4, \quad \gamma_1 = 2, \quad \gamma_2 = 1.$$

Consider the one-step-ahead linear forecast using two lags:

$$\hat{Y}_{t+1|t} = \alpha_1 Y_t + \alpha_2 Y_{t-1}.$$

- (b) [6 marks] Find the optimal forecast coefficients  $\alpha_1$  and  $\alpha_2$ .
  - (c) [4 marks] Compute the mean squared forecast error and hence obtain the coefficient of determination  $R^2 = 1 - \text{MSE}/\gamma_0$ .
15. Consider the time series  $\{2, 2, 3, 1, 2\}$  and the AR(1) model

$$X_t = \phi X_{t-1} + \varepsilon_t, \quad \varepsilon_t \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2).$$

- (a) [4 marks] Conditioning on  $X_1$ , the conditional log-likelihood is given by

$$\ell(\phi, \sigma^2) = -\frac{T-1}{2} \log(2\pi) - \frac{T-1}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{t=2}^T (X_t - \phi X_{t-1})^2.$$

Using the above expression, Prove that the maximum likelihood estimators are

$$\hat{\phi} = \frac{\sum_{t=2}^T X_t X_{t-1}}{\sum_{t=2}^T X_{t-1}^2}, \quad \hat{\sigma}^2 = \frac{1}{T-1} \sum_{t=2}^T (X_t - \hat{\phi} X_{t-1})^2,$$

and hence show that the maximised log-likelihood reduces to

$$\ell(\hat{\phi}, \hat{\sigma}^2) = -\frac{T-1}{2} \log(\hat{\sigma}^2) - \frac{T-1}{2} (1 + \log 2\pi).$$

- (b) [3 marks] Compute  $\hat{\phi}$  and  $\hat{\sigma}^2$  for the given data.
- (c) [3 marks] Prove that for this model the AIC can be written as

$$\text{AIC} = 4 \log\left(\frac{11\pi}{4}\right) + 8,$$

and evaluate this expression numerically.

16. We know that that the autocovariance of a MA(1) process vanishes above lag 1. In this problem, I will help you prove the converse for invertible covariance stationary processes. Suppose  $(X_t)$  is invertible covariance stationary with autocovariance  $\gamma(h) = 0$  for all  $|h| \geq 2$ .

- (a) [4 marks] Show that

$$g_X(e^{-i\omega}) = \gamma(0) + 2\gamma(1) \cos \omega.$$

Hint: Use  $e^{i\omega} = \cos \omega + i \sin \omega$ .

- (b) [4 marks] Equating the above expression to autocovariance generating function of MA(1)

$$g_X(e^{-i\omega}) = \sigma^2 |1 + \theta e^{-i\omega}|^2$$

derive a **unique** value for  $\sigma^2, \theta$ .

- (c) [2 marks] Assuming the uniqueness of Fourier series representation, conclude that  $(X_t)$  is MA(1) process.

17. We will work out the spurious regression between two independent random walks. Consider two independent I(1) processes:

$$y_t = y_{t-1} + u_t, \quad x_t = x_{t-1} + v_t, \quad t = 1, \dots, T,$$

where  $u_t \stackrel{\text{iid}}{\sim} (0, \sigma^2)$ ,  $v_t \stackrel{\text{iid}}{\sim} (0, \sigma^2)$ ,  $\{u_t\}$  and  $\{v_t\}$  are mutually independent, and  $y_0 = x_0 = 0$ . Consider the OLS regression *without intercept*:  $y_t = \beta x_t + \varepsilon_t$ , with coefficient of determination:

$$R^2 = \frac{\left(\sum_{t=1}^T y_t x_t\right)^2}{\sum_{t=1}^T y_t^2 \cdot \sum_{t=1}^T x_t^2}.$$

- (a) [3 marks] By the functional central limit theorem (FCLT):

$$\frac{1}{\sigma\sqrt{T}} y_{\lfloor rT \rfloor} \xrightarrow{d} W_1(r), \quad \frac{1}{\sigma\sqrt{T}} x_{\lfloor rT \rfloor} \xrightarrow{d} W_2(r),$$

where  $W_1, W_2$  are independent standard Brownian motions on  $[0, 1]$ . Using the Riemann sum approximation, show that:

$$\frac{1}{T^2} \sum_{t=1}^T y_t^2 \xrightarrow{d} \sigma^2 \int_0^1 W_1(r)^2 dr.$$

- (b) [4 marks] Using the continuous mapping theorem, show that:

$$R^2 \xrightarrow{d} \frac{\left(\int_0^1 W_1(r) W_2(r) dr\right)^2}{\int_0^1 W_1(r)^2 dr \cdot \int_0^1 W_2(r)^2 dr}.$$

- (c) [3 marks] It turns out that the limiting distribution has  $E[R^2] \approx 0.44$ . Explain why this invalidates  $R^2$  as a measure of fit for I(1) regressions, and state one remedy.